

1. (Preliminary Work)      DEF    The Gamma Function is defined as  $\Gamma(t) = \int_0^\infty y^{t-1} e^{-y} dy$     for  $t > 0$ .

(a). Use integration by parts once to rewrite the Gamma Function.    [Remember you are integrating with respect to  $y$ .]

$$u = y^{t-1} \qquad dv = e^{-y} dy$$

$$du = \underline{\hspace{2cm}} dy \qquad v = \underline{\hspace{2cm}}$$

$$\Gamma(t) = \int_0^\infty y^{t-1} e^{-y} dy = \underline{\hspace{2cm}} \Big|_0^\infty + \int_0^\infty (t-1)y^{t-2} e^{-y} dy$$

$$= \lim_{y \rightarrow \infty} (-y^{t-1} e^{-y}) + \underline{\hspace{2cm}} + (t-1) \int_0^\infty \underline{\hspace{2cm}}$$

The limit goes to  $\underline{\hspace{2cm}}$  since  
 exponentials decay much faster than any power grows.

$$= 0 + 0 + (t-1) \int_0^\infty y^{t-2} e^{-y} dy$$

$$= (t-1) \underbrace{\int_0^\infty y^{t-2} e^{-y} dy}_{\Gamma(\quad)}$$

$$= (t-1)\Gamma(t-1)$$

i.e.  $\Gamma(t) = (t-1)\Gamma(t-1)$

(b). Use the formula from part (a) to find and simplify  $\Gamma(5)$     (i.e.  $t = 5$ ).

$$\Gamma(5) = 4\Gamma(4) = 4(3\Gamma(3)) = 4 \cdot 3(\underline{\hspace{2cm}}) = 4 \cdot 3 \cdot 2(\underline{\hspace{2cm}}) = 4 \cdot 3 \cdot 2 \cdot 1 \cdot \Gamma(1) = \underline{\hspace{2cm}} \cdot \Gamma(1)$$

(c). Use the result of part(b) to guess the formula for  $t = n$  a positive integer:  $\Gamma(n) = \underline{\hspace{2cm}}$

(d). Evaluate the integral to determine  $\Gamma(1)$ .

$$\Gamma(1) = \int_0^\infty y^0 e^{-y} dy = \int_0^\infty e^{-y} dy =$$

(e). Combine the results of parts (c) and (d):

$\Gamma(n) =$

We'll use this formula later.

Poisson Distribution with mean  $\lambda$  (i.e.  $\lambda$  occurrences per unit interval):

$$f(x) = \frac{\lambda^x e^{-\lambda}}{x!}, \quad x = 0, 1, \dots$$

$\Rightarrow$

Exponential Distribution with parameter  $\theta = \frac{1}{\lambda}$  (i.e. mean time until the 1<sup>st</sup> occurrence):

$$f(x) = \frac{1}{\theta} e^{-x/\theta}, \quad x \geq 0$$

Suppose we now want to consider the random variable  $X$  that represents the waiting time until the  $\alpha^{\text{th}}$  occurrence.

**2.** Derive the c.d.f for  $X$ .

$$F(x) = P(X \leq x)$$

$$= 1 - P(X > x) \quad \text{i.e. The } \underline{\hspace{2cm}} \text{ occurrence happens after time } \underline{\hspace{2cm}} .$$

$$= 1 - P(\text{less than } \underline{\hspace{2cm}} \text{ occurrences in the interval } \underline{\hspace{2cm}})$$

$$= 1 - P(Y < \alpha) \quad \text{where } Y \text{ is the number of occurrences in the interval } [0, x].$$

So  $Y$  is  $\underline{\hspace{2cm}}$  with parameter  $\underline{\hspace{2cm}}$  .

$$= 1 - [P(Y = 0) + P(Y = 1) + P(Y = 2) + \dots + P(Y = \underline{\hspace{2cm}})]$$

$$= 1 - \sum_{k=0}^{\alpha-1} \frac{(\lambda x)^k e^{-\lambda x}}{k!}$$

$$\text{i.e. } F(x) = 1 - \sum_{k=0}^{\alpha-1} \frac{(\lambda x)^k e^{-\lambda x}}{k!}$$

3. Use the c.d.f  $F(x) = 1 - \sum_{k=0}^{\alpha-1} \frac{(\lambda x)^k e^{-\lambda x}}{k!}$  to derive the p.d.f for  $X$ . Recall  $f(x) = F'(x)$

Then since  $F(x) = 1 - e^{-\lambda x} - e^{-\lambda x} \sum_{k=1}^{\alpha-1} \frac{(\lambda x)^k}{k!}$  Separated out the \_\_\_\_\_ term and

factored out  $e^{-\lambda x}$  from the sum since it does not depend on \_\_\_\_\_ .

$$f(x) = F'(x) = 0 - e^{-\lambda x} \cdot (\text{_____}) - \left( \text{_____} \sum_{k=1}^{\alpha-1} \frac{k(\lambda x)^{k-1} \cdot \lambda}{k!} + \sum_{k=1}^{\alpha-1} \frac{(\lambda x)^k}{k!} \cdot e^{-\lambda x} \cdot (-\lambda) \right) \quad [\text{Prod. Rule}]$$

$$= \lambda e^{-\lambda x} - \left( \lambda e^{-\lambda x} \sum_{k=1}^{\alpha-1} \frac{k(\lambda x)^{k-1}}{k!} - \lambda e^{-\lambda x} \sum_{k=1}^{\alpha-1} \frac{(\lambda x)^k}{k!} \right)$$

$$= \lambda e^{-\lambda x} - \lambda e^{-\lambda x} \sum_{k=1}^{\alpha-1} \left[ \frac{k(\lambda x)^{k-1}}{k!} - \frac{(\lambda x)^k}{k!} \right] \quad \text{but } \frac{k}{k!} = \frac{k}{k \cdot (k-1)!} = \frac{1}{(k-1)!}$$

$$= \lambda e^{-\lambda x} - \lambda e^{-\lambda x} \sum_{k=1}^{\alpha-1} \left[ \frac{(\lambda x)^{k-1}}{(k-1)!} - \frac{(\lambda x)^k}{k!} \right]$$

$$= \lambda e^{-\lambda x} - \lambda e^{-\lambda x} \left[ \left( \frac{(\lambda x)^0}{0!} - \frac{(\lambda x)^1}{1!} \right) + \left( \frac{(\lambda x)^1}{1!} - \frac{(\lambda x)^2}{2!} \right) + \dots \right]$$

$$\dots + \left( \frac{(\lambda x)^{\alpha-2}}{(\alpha-2)!} - \frac{(\lambda x)^{\alpha-1}}{(\alpha-1)!} \right)$$

$$= \lambda e^{-\lambda x} - \lambda e^{-\lambda x} \left[ 1 - \frac{(\lambda x)^{\alpha-1}}{(\alpha-1)!} \right] \quad \text{Cancel out terms above.}$$

$$= \lambda e^{-\lambda x} - \lambda e^{-\lambda x} + \lambda e^{-\lambda x} \frac{(\lambda x)^{\alpha-1}}{(\alpha-1)!}$$

$$= \lambda e^{-\lambda x} \frac{(\lambda x)^{\alpha-1}}{(\alpha-1)!}$$

i.e. The p.d.f. of  $X$  is given by  $f(x) = \lambda e^{-\lambda x} \frac{(\lambda x)^{\alpha-1}}{(\alpha-1)!}$ .

4. Let  $\lambda = \frac{1}{\theta}$  and also use the formula at the bottom of page 1 for the Gamma Function rewrite the p.d.f  $f(x)$ .

$$f(x) = \frac{\lambda e^{-\lambda x} (\lambda x)^{\alpha-1}}{(\alpha-1)!}$$

$$= \frac{\frac{1}{\theta} e^{-x/\theta} \left(\frac{1}{\theta} x\right)^{\alpha-1}}{(\alpha-1)!}$$

$$= \frac{\frac{1}{\theta} e^{-x/\theta} \left(\frac{1}{\theta}\right)^{\alpha-1} x^{\alpha-1}}{\Gamma(\alpha)}$$

$$\text{But } \frac{1}{\theta} \cdot \left(\frac{1}{\theta}\right)^{\alpha-1} = \underline{\hspace{2cm}}$$

$$= \frac{e^{-x/\theta} \left(\frac{1}{\theta}\right)^{\alpha} x^{\alpha-1}}{\Gamma(\alpha)}$$

i.e.

$f(x) = \frac{1}{\Gamma(\alpha)\theta^{\alpha}} x^{\alpha-1} e^{-x/\theta}, \quad 0 \leq x < \infty$	Gamma Distribution with parameters $\alpha$ and $\theta = \frac{1}{\lambda}$
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gives the waiting time until the  $\alpha^{\text{th}}$  occurrence in a Poisson Process.

Denoted  $\Gamma(\alpha, \theta)$ .

Note:

- Derived for  $\alpha$  an integer, but also has applications for non-integer  $\alpha$ .
- $\alpha$  is the shape parameter
- $\theta$  is the scale parameter

[Dr. Crawford will show typical sketches of  $f(x)$ .]

For the Gamma Distribution,

You will derive these formulas in the homework

$$\mu = \alpha\theta \qquad \sigma^2 = \alpha\theta^2$$

$$M(t) = \frac{1}{(1 - \theta t)^\alpha}, \quad t < \frac{1}{\theta}$$

Ex The number of times that a network computer fails has a Poisson distribution with a mean of once per month (30 days).

(a). Find the probability that it will be at least 3 months until the 2<sup>nd</sup> failure. (Use months)

(b). What if you used days (30 days = 1 month)

(c). Find  $\mu$  and  $\sigma^2$

Homework: Section 3.2, p. 109: #7, 9, 10, 12, 13, 16, 17[Hint: Two-partner with binomial distribution], 19[Set up the integrals for E(Profit) and explain how you would solve it – but do not actually solve it.]